# **Some Estimates in Variational Quantum Monte Carlo**

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July 2009

- Variational Monte Carlo
- $\bullet$  'Standard' VMC is  $P=\psi^2$  , and failure of CLT is an artifact of this
- Conditions and expressions for Normally distributed estimates in general sampling
- $\bullet$  Efficient sampling Monte Carlo implemented not sampling from  $P=\psi^2$
- Estimates for energies, excitation energies, ionisation energies, transition moments
- Results for first row atoms and some molecules

#### **VMC** and General Sampling

Sample with  $P=\psi^2/w$  and construct an estimate:

$$\overline{\mu} = \frac{\sum w_i E_L(\mathbf{R}_i)}{\sum w_i}$$

• If both variances exist, Fieller's theorem tells us this is a sample from a Normal distribution with:

$$\mu = \frac{\int \psi^2 E_L d\mathbf{R}}{\int \psi^2 d\mathbf{R}} \quad , \quad \sigma^2 = \frac{1}{r} \frac{\int \psi^2 / w d\mathbf{R} \int w \psi^2 (E_L - \mu)^2 d\mathbf{R}}{\left[\int \psi^2 d\mathbf{R}\right]^2}$$

We can estimate the variance:

$$\overline{\sigma}^2 = \frac{r}{r-1} \frac{\sum w_i^2 \left( E_L(\mathbf{R}_i) - \overline{\mu} \right)^2}{\left( \sum w_i \right)^2}$$

- ullet  $\overline{\sigma}^2 
  eq (\text{sample variance})/r$
- These equations do not follow from the usual (univariate) Central Limit Theorem
- ullet Zero Variance Principle is still valid for exact  $\psi \Rightarrow \overline{\sigma} = 0$
- ightarrow The error is **controlled** if the bivariate CLT is **valid** and  $\langle w \rangle \neq 0$

Trail JR, Phys. Rev. E. 77, 016703,016704 (2008)

#### When is it Normal?

For standard sampling  $P=\psi^2$  (w=1)

- ullet Normal for standard sampling and total energy ( $P \propto 1/x^4$ )
- Not Normal for standard sampling used with correlated sampling, forces, and many other estimates
- ullet Different choices of P (equivalently w) are possible
- ullet Changes computational cost: flops for evaluating P
- Changes distribution of random errors
- Failure of CLT from singularities in averaged quantities on the nodal surface

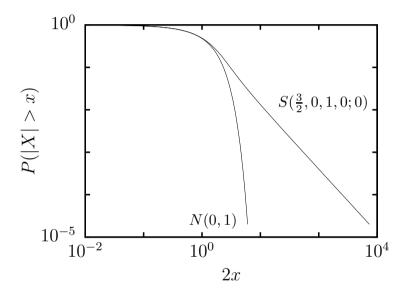
For some P (equivalently w) the variance is infinite, and CLT is invalid

ightarrow Goal is to improve efficiency and reinstate the CLT where it is invalid for standard sampling

#### Why is Normality so important?

We want Normally distributed estimates, from a general form of the Central Limit Theorem

If the distribution the estimates are drawn from is *not* normal then it is a Stable Law:



- Normal distribution and an example Stable law
- ullet Probability that a sample fall outside of central interval size 2x
- Width parameter is representative of error for Normal, not for Stable
- Width parameter is estimateable for Normal sample standard error
- Width parameter is not estimateable for Stable sample standard error is unrelated

### **Efficient sampling in VMC**

- ullet Draw position vectors from  $P=|D_1|^2+|D_2|^2$  a, no Jastrow, Backflow
- → Zero on coalescence planes only, non-zero on rest of nodal surface
- ullet Perform Metropolis accept/reject with P, and use  $w=\psi^2/P$
- Using r samples, the Bivariate CLT, and Fiellers theorem provides the Normal estimate

$$\overline{E}_{tot} = \frac{\sum w_i E_L(\mathbf{R}_i)}{\sum w_i} \quad , \quad \overline{\sigma}^2 = \frac{r}{r-1} \frac{\sum w_i^2 \left( E_L(\mathbf{R}_i) - \overline{E}_{tot} \right)^2}{\left( \sum w_i \right)^2}$$

ullet  $(wE_L,w)$  has no singularities and is bounded  $\Rightarrow$  all moments exist  $\Rightarrow$  distribution is Normal

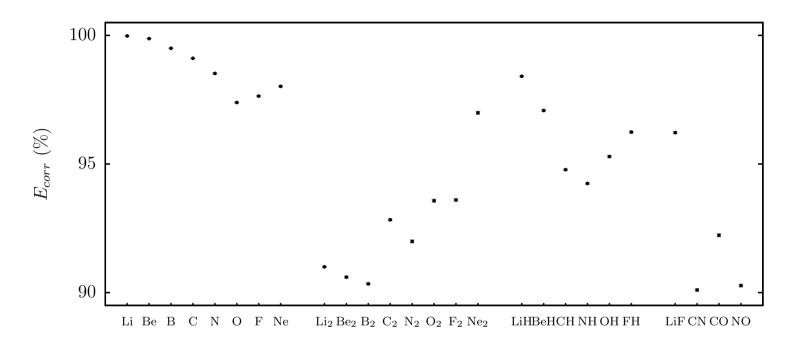
<sup>&</sup>lt;sup>a</sup>Trail JR and Maezono R, JCP (2010)

### **Efficient sampling in VMC**

Optimisation and estimation of total energy:

- All-electron
- First row atoms + some diatomic molecules
- Numerical orbitals from ATSP2K and 2DHF (MCSCF and HF)
- $\bullet$  Jastrow, Backflow, and  $5-86~\mathrm{CSFs}$
- 48 h desktop time/system
- $\bullet$  r for Final estimate:monitor estimate:optimisation r=9000:150:1

### **Efficient sampling in VMC**



- Improved results for less computational effort
- $\bullet$  For a given error  $5-35\times$  faster than standard sampling
- Surfaces for energy optimisation are Normal
- Do we need anisotropic Jastrow/Backflow?

#### **Distribution of more general estimates**

What is the distribution for more complex estimates?

### **Example: Energy differences**

$$\Delta E_{tot} = \frac{\int \psi_1^2 E_1 d\mathbf{R}}{\int \psi_1^2 d\mathbf{R}} - \frac{\int \psi_2^2 E_2 d\mathbf{R}}{\int \psi_2^2 d\mathbf{R}}$$

ullet Sample with  $P=|D_1|^2+|D_2|^2$  (no nodal surface, only coalescence planes)

Est 
$$[\Delta E_{tot}]$$
 =  $\frac{\sum w_1 E_1}{\sum w_1} - \frac{\sum w_2 E_2}{\sum w_2}$   
 =  $\frac{S_2}{S_1} - \frac{S_4}{S_3}$ 

- ullet Elements in sums  $S_n$  are correlated only when in same 'time slots'
- ⇒ Multivariate CLT is true
- $\Rightarrow$  Each  $S_n$  is Normal
- $\Rightarrow$  Correlation between  $S_n$  and  $S_m$  is linear

### Algebra for sums of random variables

What is the distribution of

$$\Delta \mathsf{E}_{tot} = rac{\mathsf{S}_2 \mathsf{S}_3 - \mathsf{S}_1 \mathsf{S}_4}{\mathsf{S}_1 \mathsf{S}_3}$$
 ,  $\mathsf{S}_n = \sum_i X_n(i)$ 

with parameters expressed in terms of the estimateable

$$\mathbb{E} \left[ \mathsf{X}_{n}(i) \right] = \mu_{n}$$

$$Var \left[ \mathsf{X}_{n}(i) \right] = C_{nn}$$

$$Cov \left[ \mathsf{X}_{m}(i), \mathsf{X}_{n}(j) \right] = C_{nm} \delta_{i,j}$$

### Algebra for sums of random variables : Sums

Adding sums of the random variables, trivial to show that:

Distribution is Normal

$$\mathbb{E}\left[\mathsf{S}_{1} + \mathsf{S}_{2}\right] = r(\mu_{1} + \mu_{2})$$

$$Var\left[\mathsf{S}_{1} + \mathsf{S}_{2}\right] = r(C_{11} + 2.C_{12} + C_{22})$$

$$Cov\left[\mathsf{S}_{1} + \mathsf{S}_{2}, \mathsf{S}_{3} + \mathsf{S}_{4}\right] = r(C_{13} + C_{14} + C_{23} + C_{24})$$

- $\rightarrow$  All the  $\mathsf{S}_1+\mathsf{S}_2$  are Normal
- ightarrow Correlation between any  $\mathsf{S}_j + \mathsf{S}_k$  and  $\mathsf{S}_l + \mathsf{S}_m$  is linear

### Algebra for sums of random variables: Quotients

Quotients of sums of the random variables, Fiellers theorem provides:

Distribution is Normal

$$\mathbb{E}\left[\frac{\mathsf{S}_{2}}{\mathsf{S}_{1}}\right] = \frac{\mu_{2}}{\mu_{1}}$$

$$Var\left[\frac{\mathsf{S}_{2}}{\mathsf{S}_{1}}\right] = \frac{1}{r}\frac{1}{\mu_{1}^{2}}\left[C_{22} - 2\frac{\mu_{2}}{\mu_{1}}C_{12} + \left(\frac{\mu_{2}}{\mu_{1}}\right)^{2}C_{11}\right]$$

• No expression for correlation of different quotients

So,

$$\frac{\mathsf{S}_1 + \mathsf{S}_2}{\mathsf{S}_3 + \mathsf{S}_4}$$

is Normal and parameters are estimateable

Products of sums of the random variables:

- ullet What is the bivariate distribution of  $(\mathsf{S}_1\mathsf{S}_2,\mathsf{S}_3\mathsf{S}_4)$  ?
- ullet Derive co-moments and compare with bivariate Normal....define  $\Delta {\sf X}_1(i) = {\sf X}_1(i) \mu_1$

$$\Delta(S_1S_2) = S_1.S_2 - \mathbb{E}[S_1.S_2] 
= \sum_{ij} \Delta X_1(i) \Delta X_2(j) + r \sum_i [\mu_1 \Delta X_2(i) + \mu_2 \Delta X_1(i)] - rC_{12}$$

ullet Co-moments are defined by  $\mu_{m,n}=\mathbb{E}\left[\Delta(\mathsf{S}_1\mathsf{S}_2)^m.\Delta(\mathsf{S}_3\mathsf{S}_4)^n
ight]$ 

$$\mu_{m,n} = \mathbb{E}\left[\left(\sum_{ij} \Delta \mathsf{X}_1(i) \Delta \mathsf{X}_2(j) + r \sum_{i} \left[\mu_1 \Delta \mathsf{X}_2(i) + \mu_2 \Delta \mathsf{X}_1(i)\right] - r C_{12}\right)^m \times \left(\sum_{ij} \Delta \mathsf{X}_3(i) \Delta \mathsf{X}_4(j) + r \sum_{i} \left(\mu_3 \Delta \mathsf{X}_4(i) + \mu_4 \Delta \mathsf{X}_3(i)\right] - r C_{34}\right)^n\right]$$

- Multiply out and count equivalent terms
- Pick out the dominant r terms.

Powers of r in each term after multiplication:

- ullet  $2^{nd}$  order part: Powers  $p_1,q_1$  from equivalence, none from prefactor
- ullet 1 order part: Powers  $p_2,q_2$  from equivalence,  $p_2,q_2$  from prefactor
- ullet 0 order part: Powers  $p_3,q_3$  from equivalence,  $p_3,q_3$  from prefactor
- Multiplication gives  $p_1 + p_2 + p_3 = m$ ,  $q_1 + q_2 + q_3 = n$
- ullet Count powers of r from prefactors [r], and from sums (r)
- $\Rightarrow$  Terms have powers of r:

$$[1]^{p_1}(r)^{p_1}[r]^{p_2}(r)^{p_2}[r]^{p_3}(1)^{p_3}.[1]^{q_1}(r)^{q_1}[r]^{q_2}(r)^{q_2}[r]^{q_3}(1)^{q_3}\mathbb{E}\left[\ldots\right]$$

But Expectations are zero if any index is unique:

$$\mathbb{E} \left[ \Delta \mathsf{X}_{1}(i).\Delta \mathsf{X}_{2}(j).\Delta \mathsf{X}_{3}(k) \cdots \right] = \mathbb{E} \left[ \Delta \mathsf{X}_{1}(i) \right].\mathbb{E} \left[ \Delta \mathsf{X}_{2}(j).\Delta \mathsf{X}_{3}(k) \cdots \right] = 0$$

Overcounting due to including zero expectations values

Count only terms with no unique indices:

$$r^{(m+n)+(p_2+q_2)/2}\mathbb{E}\left[\ldots\right]$$
 for  $(p_2+q_2)$  even  $r^{(m+n)+(p_2+q_2+1)/2}\mathbb{E}\left[\ldots\right]$  for  $(p_2+q_2)$  odd

- $\Rightarrow$  In large r limit  $p_2+q_2=m+n$  terms dominate ( $p_1=p_3=q_1=q_3=0$ )
- $\Rightarrow$  In large r limit the co-moments are same as those for *sums*:

$$\mathbb{E}\left[\Delta(\mathsf{S}_{1}\mathsf{S}_{2})^{m}.\Delta(\mathsf{S}_{3}\mathsf{S}_{4})^{n}\right] = r^{m+n}\mathbb{E}\left[\left(\sum_{i}\left[\mu_{1}\Delta\mathsf{X}_{2}(i) + \mu_{2}\Delta\mathsf{X}_{1}(i)\right]\right)^{m}\right] \times \left(\sum_{i}\left[\mu_{3}\Delta\mathsf{X}_{4}(i) + \mu_{4}\Delta\mathsf{X}_{3}(i)\right]\right)^{n}\right]$$

ullet In large r limit  $(\mathsf{S}_1\mathsf{S}_2,\mathsf{S}_3\mathsf{S}_4)$  is bivariate Normal with

$$\mathbb{E}[\mathsf{S}_1\mathsf{S}_2] = r^2\mu_1.\mu_2$$

$$Var[\mathsf{S}_1\mathsf{S}_2] = r^3(\mu_1^2C_{22} + 2\mu_1\mu_2C_{12} + \mu_2^2C_{11})$$

$$Cov[\mathsf{S}_1\mathsf{S}_2,\mathsf{S}_3\mathsf{S}_4] = r^3(\mu_1\mu_3C_{24} + \mu_1\mu_4C_{23} + \mu_2\mu_3C_{14} + \mu_2\mu_4C_{13})$$

ightarrow We have rules for obtaining the distribution of combinations of sums of random variables (that are Normal)

### **Energy differences**

Est 
$$[\Delta E_{tot}]$$
 =  $\frac{\sum w_1 E_1}{\sum w_1} - \frac{\sum w_2 E_2}{\sum w_2}$   
 =  $\frac{\mathsf{S}_2 \mathsf{S}_3 - \mathsf{S}_1 \mathsf{S}_4}{\mathsf{S}_1 \mathsf{S}_2}$ 

$$\overline{\Delta E}_{tot} = \frac{\overline{\mu}_2}{\overline{\mu}_1} - \frac{\overline{\mu}_4}{\overline{\mu}_3} 
\frac{r-1}{r} \overline{\sigma}^2 = \frac{\sum w_1^2 (E_1 - \overline{\mu}_1)^2}{\left[\sum w_1\right]^2} - 2 \frac{\sum w_1 w_2 (E_1 - \overline{\mu}_1) (E_2 - \overline{\mu}_2)}{\left[\sum w_1\right] \left[\sum w_2\right]} + \frac{\sum w_2^2 (E_2 - \overline{\mu}_2)^2}{\left[\sum w_2\right]^2}$$

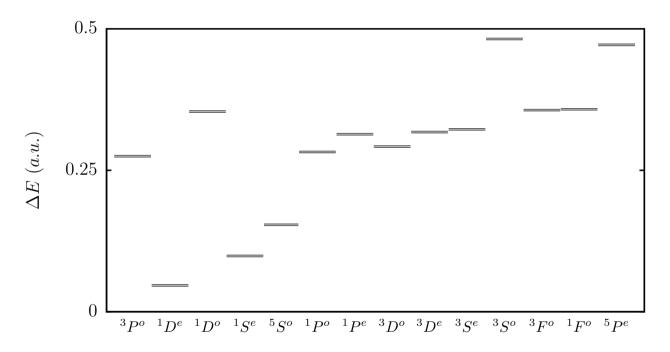
Sufficient conditions for this estimate to be Normal are

- All moments exist
- All means in the denominator are non-zero
- At least one mean in each product on numerator is non-zero

### **Carbon atom excitation energies**

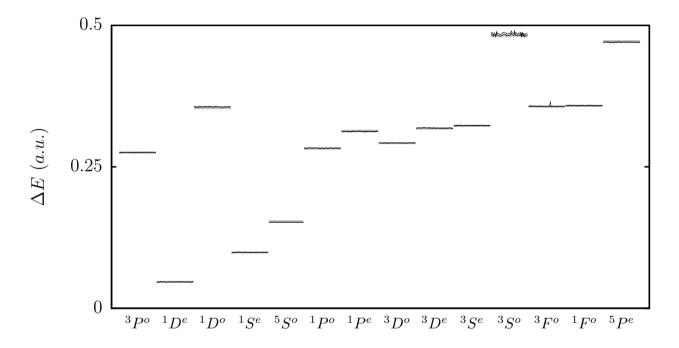
- $\bullet$  Calculation as for GS (Multideterminant and numerical orbitals from MCSCF  $\sim 300$  parameters)
- $E_{tot}$   $\begin{bmatrix} 2S+1L \end{bmatrix} E_{tot} \begin{bmatrix} 3P \end{bmatrix}$
- Lowest two eigenstates for each Term
- Inversion symmetry conserved by Jastrow and Backflow
- Term approximately conserved by introduction of Jastrow and Backflow
- Energy minimisation approximately valid for lowest two energies for each Term
- Sample with  $P = D_1 \left[ {^{2S+1}L} \right]^2 + D_2 \left[ {^{2S+1}L} \right]^2 + D_1 \left[ {^3P} \right]^2 + D_2 \left[ {^3P} \right]^2$

## **Carbon atom excitation energies**



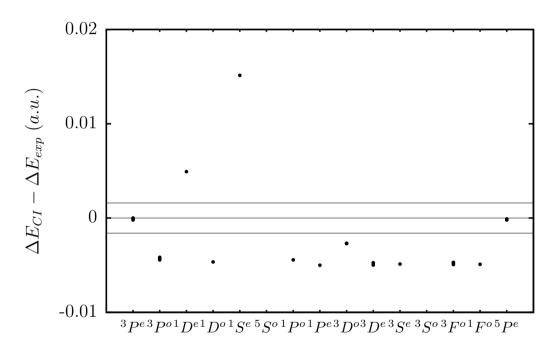
- $\bullet$  C  $^{2S+1}L \rightarrow$  C  $^3P$
- ullet Grey: experimental Spectroscopic values  $\pm$  'chemical accuracy'

## **Carbon atom excitation energies**



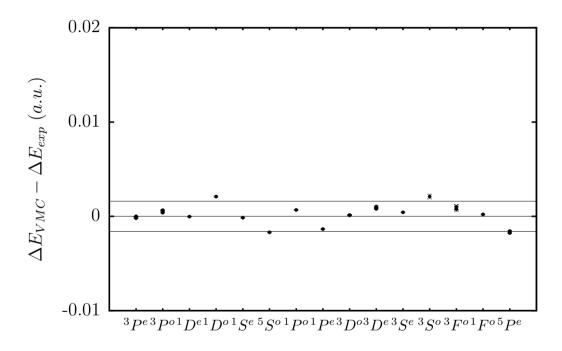
- ullet C  $^{2S+1}L o$  C  $^3P$
- ullet Grey: experimental Spectroscopic values  $\pm$  'chemical accuracy'
- 48 energy difference estimates for each excitation

# **Error in Carbon atom excitation energies**



- CI results used for allocating transitions to lines
- CI AS and orbitals chosen empirically to reproduce spectrosopy

## **Error in Carbon atom excitation energies**



- Chemical accuracy from VMC + efficient sampling
- NOT spectroscopic accuracy
- Estimate of difference not difference of estimates
- $\bullet$  Correlation reduces error by 10-70%

### **Ionization energies**

• How do we deal with changes in electron number?

$$E_{ion} = E_n - E_{n-1}$$

$$= \frac{\int \psi_1^2 E_1 d\mathbf{R}_n}{\int \psi_1^2 d\mathbf{R}_n} - \frac{\int \psi_2^2 E_2 d\mathbf{R}_{n-1}}{\int \psi_2^2 d\mathbf{R}_{n-1}}$$

- Sampled  $P_n(\mathbf{R}_n) = D_1(n)^2 + D_2(n)^2$
- ullet For estimating  $E_{n-1}$  ignore *one* 3d sample vector, so  ${f R}_n o {f R}_{n-1}$
- ullet Distribution of  $(w_2E_2,w_2)$  given by integrating  $P_n$  analytically

$$P_{n-1} = \int D_1(n)^2 + D_2(n)^2 d^3 \mathbf{r}_n$$

$$= \int [\phi_1(\mathbf{r}_n) \cdot C_1(1, n) + \phi_2(\mathbf{r}_n) \cdot C_1(2, n) + \dots]^2 + [\phi_1(\mathbf{r}_n) \cdot C_2(1, n) + \phi_2(\mathbf{r}_n) \cdot C_2(2, n) + \dots]^2 d^3 \mathbf{r}_n$$

$$= [C_1(1, n)^2 + C_1(2, n)^2 + \dots] + [C_2(1, n)^2 + C_2(2, n)^2 + \dots]$$

### **lonization energies**

ullet Provides distributions of  $(w_1E_1,w_1,w_2E_2,w_2)$  in terms of weights

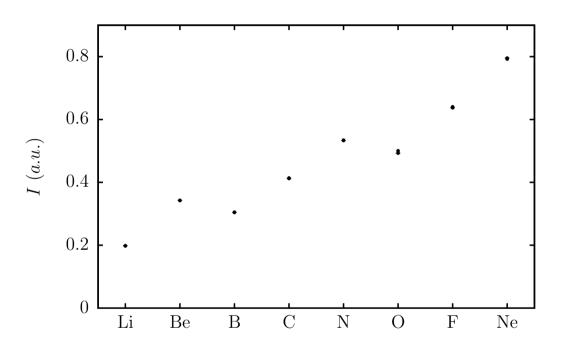
$$w_1(\mathbf{R}_n) = \psi_1^2(\mathbf{R}_n)/P_n(\mathbf{R}_n)$$
  
$$w_2(\mathbf{R}_{n-1}) = \psi_2^2(\mathbf{R}_{n-1})/P_{n-1}(\mathbf{R}_{n-1})$$

- ullet  $P_n$  is zero on coalescence planes only
- $\bullet$   $P_{n-1}$  is zero on coalescence planes only
- ullet All sums are Normal and linearly correlated o Normal estimate

### **Ionization energies**

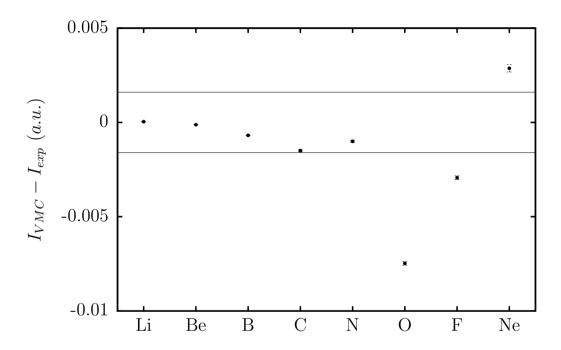
- First row neutral atoms and ions
- Orbitals from same source as before
- Jastrow/Backflow and computational cost as before
- Optimise neutral atom and ion seperately, with energy minimisation
- Estimate energy difference/error as for excitation energies

# **Ionization energies**



- Experimental ionization energies
- VMC estimated energy difference

# **Error in Ionization energies**



- Close but *not* chemical accuracy
- Maybe we need T/Q excitations?
- $\bullet$  Correlation reduces error by 0-40%

Spectroscopic line widths characterised by transition dipole moments

$$\Delta_{\omega} = \frac{3}{2} \Delta E \sum_{m} |\langle \psi_0 | \mathbf{R} | \psi_m \rangle|^2$$

with sum over total angular momentum eigenstates.

Not done yet ... start with estimates for transition dipole moments:

$$t_{12} = |\langle \psi_1 | \sum_i \mathbf{r}_i | \psi_2 \rangle|^2$$

- Sample using  $P = D_1(1)^2 + D_2(1)^2 + D_1(2)^2 + D_2(2)^2$
- Get an error estimate from random variable algebra
- Not zero variance...

Estimate:

$$t_{12} = |\langle \psi_1 | \sum_{\mathbf{r}_i | \psi_2 \rangle|^2}$$

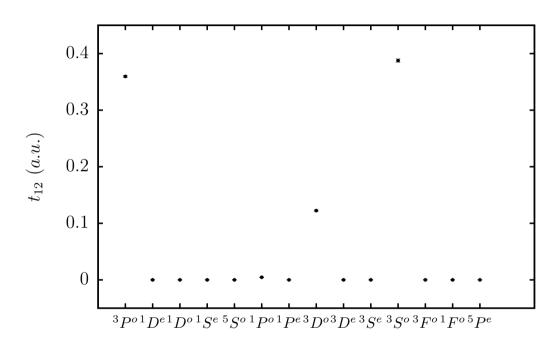
$$= \frac{\left[ \int \psi_1 \psi_2(x_1 + \dots + x_n) d\mathbf{R} \right]^2 + \left[ \int \psi_1 \psi_2(y_1 + \dots + y_n) d\mathbf{R} \right]^2 + \left[ \int \psi_1 \psi_2(z_1 + \dots + z_n) d\mathbf{R} \right]^2}{\int \psi_1^2 d\mathbf{R} \int \psi_2^2 d\mathbf{R}}$$

- Define weights  $w_1 = \psi_1^2/P$ ,  $w_2 = \psi_2^2/P$
- Analyse as before ...

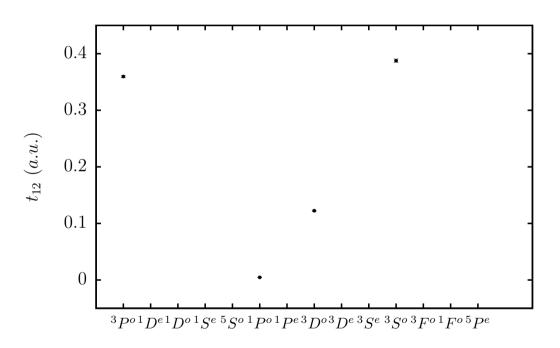
$$\mathsf{Est}\left[t_{12}\right] = \frac{\mathsf{S}_{1}^{2} + \mathsf{S}_{2}^{2} + \mathsf{S}_{3}^{2}}{\mathsf{S}_{4}.\mathsf{S}_{5}}$$

• Normally distributed with estimateable mean and variance:

• Replace C's and  $\mu$ 's with unbiased estimates



• Many are non-normal and zero from symmetry considerations...



- ... so we drop them
- $\bullet$  Random error  $\sim 1\%$

#### **Conclusions**

- Normal errors can be reintroduced
- More computationally efficient than standard sampling
- Optimisation is on a Normal surface, unlike standard sampling
- Distribution of random error can be derived for general estimates and sampling

Normal and efficient estimates implemented for:

- Total energies
- Energy differences and Ionization energies
- Transition moments

### Next?

- Spectroscopic line widths
- Electron affinities
- Optimise orthogonalised trial wavefunctions for more excited states
- Normal force estimates and geometery optimisation surfaces
- Generalised DMC